

# Statistical testing of the covariance matrix rank in multidimensional neuronal models: non-asymptotic case

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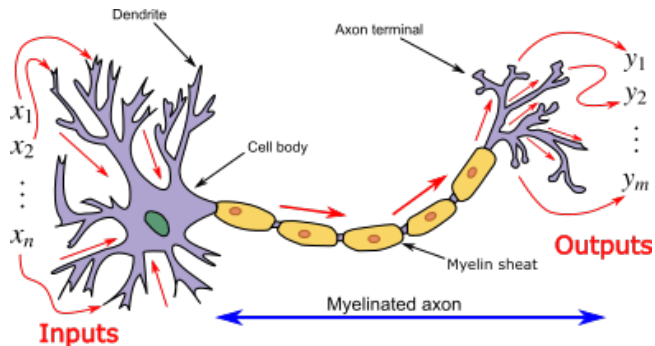
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# Motivation



# Example 1: Hodgkin-Huxley model

Conductance-based model of **action potential in neurons**:

$$\begin{cases} I &= C_m \frac{dV_m}{dt} + \bar{g}_K n^4 (V_m - V_K) + \bar{g}_{Na} m^3 h (V_m - V_{Na}) + \bar{g}_l (V_m - V_l) \\ \frac{dn}{dt} &= \alpha_n (V_m) (1 - n) - \beta_n (V_m) n \\ \frac{dm}{dt} &= \alpha_m (V_m) (1 - m) - \beta_m (V_m) m \\ \frac{dh}{dt} &= \alpha_h (V_m) (1 - h) - \beta_h (V_m) h \end{cases}$$

- ▶  $I$  – membrane potential
- ▶  $n, m, h$  – quantities between 0 and 1 that are associated with potassium channel activation, sodium channel activation, and sodium channel inactivation.

**References:** Hodgkin and Huxley (1952) – *1963 Nobel Prize in Physiology or Medicine*,

**Modifications:** Fitzhugh (1961), Morris and Lecar (1981)

## Example 2: Jansen and Rit Neural Mass model

Convolution-based model of a **neuronal population with excitatory and inhibitory subpopulations**:

$$\begin{cases} dQ(t) = \nabla_P H(Q, P) dt, \\ dP(t) = (-\nabla H(Q, P) - 2\Gamma P + G(t, Q)) dt + \Sigma(t) dW_t, \end{cases}$$

- ▶  $Q = (X_0, X_1, X_2) \in \mathbb{R}^3$ ,  $P = (X_3, X_4, X_5) \in \mathbb{R}^3$
- ▶  $\Gamma = \text{diag}[a, a, b] \in \mathbb{R}^{3 \times 3}$  is a damping part,
- ▶  $\Sigma(t) = \text{diag}[\sigma_3(t), \sigma_4(t), \sigma_5(t)] \in \mathbb{R}^{3 \times 3}$  is a diffusion part,
- ▶  $G(t, Q)$  is a nonlinear displacement term.
- ▶ Diffusion components are of **different order** (e.g.  $\sigma_3(t), \sigma_5(t) \ll \sigma_4(t)$ )!

**References:** Jansen and Rit (1995), Ableidinger et al. (2017)

# Example 2: Jansen and Rit Neural Mass model

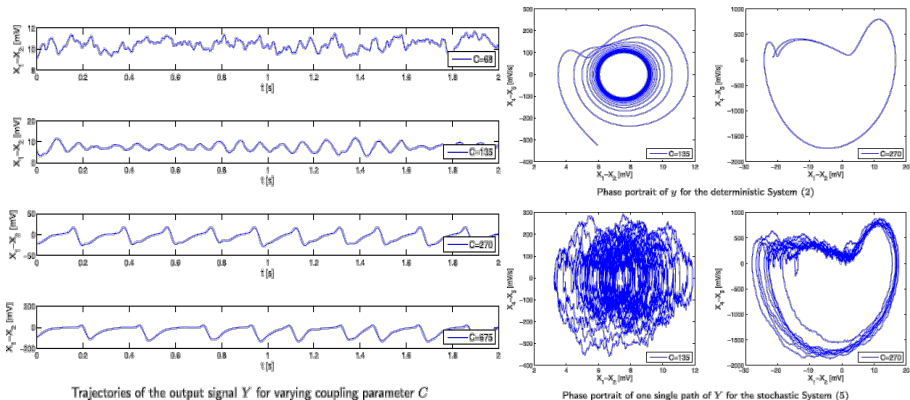


Figure: Source: Ableidinger et al. (2017)

## Example 3: Diffusion approximation of a Hawkes process

**Hawkes process** (point process with memory), describing the action potentials in a population of neurons, can be approximated by a stochastic diffusion:

$$dX_t = b_t dt + \sigma_t dW_t.$$

- ▶  $\kappa$  populations (=  $\kappa$  rough variables)
- ▶  $\sum_{i=1}^{\kappa} \eta_i$  memory variables (*smooth* variables)
- ▶  $N$  neurons

$$b_t = \begin{pmatrix} -\nu_1 x^1 + x^2 \\ -\nu_1 x^2 + x^3 \\ \vdots \\ -\nu_1 x^{\eta_1+1} + c_1 f_2(x^{\eta_1+2}) \\ -\nu_2 x^{\eta_1+2} + z^{\eta_1+3} \\ \vdots \\ -\nu_n x^\kappa + c_n f_1(x^1) \end{pmatrix}, \quad \sigma_t = \frac{1}{\sqrt{N}} \begin{pmatrix} 0 & 0 \\ \vdots & \vdots \\ 0 & \frac{c_1}{\sqrt{\rho_2}} \sqrt{f_2(x^{\eta_1+2})} \\ 0 & 0 \\ \vdots & \vdots \\ \frac{c_n}{\sqrt{\rho_1}} \sqrt{f_1(x^1)} & 0 \end{pmatrix},$$

**Reference:** Ditlevsen and Löcherbach (2017), Chevallier (2017)

## Example 3: Diffusion approximation of a Hawkes process

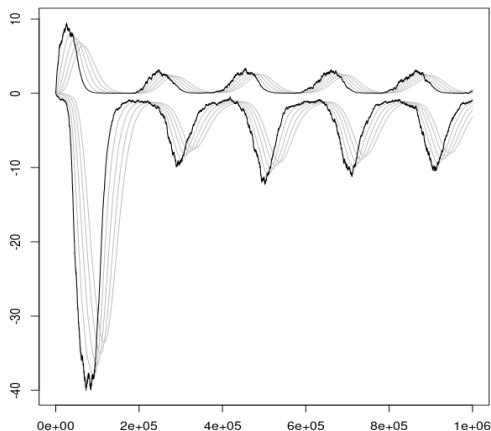


Figure: Diffusion approximation of Hawkes process describing inhibitory and excitatory neuron population (20 neurons in each population)

# Where to put noise?

## Main challenges:

- ▶ Highly non-linear systems
- ▶ Computational cost
- ▶ Measurements inaccuracy

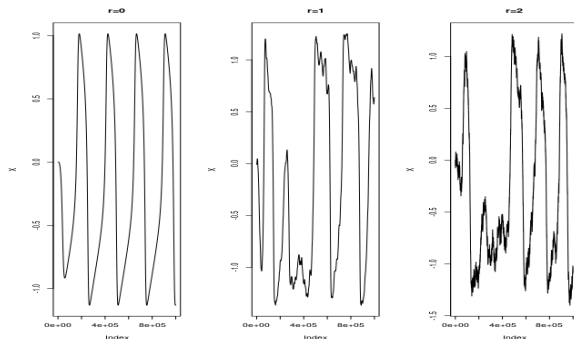


Figure: Membrane potential simulated with a FitzHugh-Nagumo model: deterministic, noisy channels, elliptic system

**References:** see Tuckwell (2005) for a general overview of neuronal models



# Formalization

Given:

Discrete observations of the  $d$ -dimensional process **with a fixed time step**  $\Delta$

$$dX_t = b_t dt + \sigma_t dW_t, \quad t \in [0, T], \quad (1)$$

$$b_t \in \mathbb{R}^d, \sigma_t \in \mathbb{R}^{d \times q}.$$

Goal:

Propose a test

$$H_0 : \text{rank}(\Sigma) = r_0$$

$$H_1 : \text{rank}(\Sigma) \neq r_0,$$

where  $\Sigma = \sigma_t \sigma_t^T$ . If  $\sigma_t$  is not constant, we search  $\sup_t \text{rank}(\Sigma)$  instead.

# Random perturbation approach: asymptotic setting

**Main references:** Jacod and Podolskij (2013) (see also Jacod et al. (2008))

Given a  $d$ -dimensional diffusion process  $X$ , consider 2 new processes:

$$\tilde{X}_t^{(k)} = X_t + \sqrt{k\Delta}\tilde{\sigma} \tilde{W}_t,$$

where  $k = 1, 2$ , and  $\tilde{\sigma}$  is such that  $\tilde{\sigma}\tilde{\sigma}^T =: \tilde{\Sigma}$  is a non-random matrix of full rank.

$$S_T^k = 2d\Delta \sum_{i=0}^{N-1} \det \begin{pmatrix} \frac{\tilde{X}_{2id+k}^{1,(k)} - \tilde{X}_{2id}^{1,(k)}}{\sqrt{k\Delta}} & \frac{\tilde{X}_{2id+2k}^{1,(k)} - \tilde{X}_{2id+k}^{1,(k)}}{\sqrt{k\Delta}} & \cdots & \frac{\tilde{X}_{2id+2d}^{1,(k)} - \tilde{X}_{2id+kd-k}^{1,(k)}}{\sqrt{k\Delta}} \\ \frac{\tilde{X}_{2id+k}^{2,(k)} - \tilde{X}_{2id}^{2,(k)}}{\sqrt{k\Delta}} & \frac{\tilde{X}_{2id+2k}^{2,(k)} - \tilde{X}_{2id+k}^{2,(k)}}{\sqrt{k\Delta}} & \cdots & \frac{\tilde{X}_{2id+2d}^{2,(k)} - \tilde{X}_{2id+kd-k}^{2,(k)}}{\sqrt{k\Delta}} \\ \cdots & \cdots & \cdots & \cdots \\ \frac{\tilde{X}_{2id+k}^{d,(k)} - \tilde{X}_{2id}^{d,(k)}}{\sqrt{k\Delta}} & \frac{\tilde{X}_{2id+2k}^{d,(k)} - \tilde{X}_{2id+k}^{d,(k)}}{\sqrt{k\Delta}} & \cdots & \frac{\tilde{X}_{2id+2d}^{d,(k)} - \tilde{X}_{2id+kd-k}^{d,(k)}}{\sqrt{k\Delta}} \end{pmatrix}^2$$

# Random perturbation approach: asymptotic setting

**Define:**

$$\hat{R}_{T,\Delta} = d - \frac{\log \frac{S_T^2}{S_T^1}}{\log 2}$$

$$V_{T,\Delta} := \text{Var} \left[ \hat{R}_{T,\Delta} \right] = \frac{\left( \frac{E[S_T^1]}{E[S_T^2]} \right)^2 \text{Var}[S_T^2] - 2 \frac{E[S_T^1]}{E[S_T^2]} \text{Cov}[S_T^1, S_T^2] + \text{Var}[S_T^1]}{(E[S_T^1] \log 2)^2}.$$

Jacod and Podolskij (2013)

$$\frac{\hat{R}_{T,\Delta} - r_0}{\sqrt{\Delta V_{T,\Delta}}} \xrightarrow{\mathcal{L}} \mathcal{N}(0, 1), \text{ as } \Delta \rightarrow 0$$

# How does it work: toy 1d example

Take the process:

$$dX_t = a dt + \sigma dW_t$$

Add a random perturbation:

$$\tilde{X}_t^{(1)} = a dt + \sigma dW_t + \sqrt{\Delta} \tilde{\sigma} \tilde{W}_t,$$

$$\tilde{X}_t^{(2)} = a dt + \sigma dW_t + \sqrt{2\Delta} \tilde{\sigma} \tilde{W}_t$$

Using the first-order approximation, compute:

$$\mathbb{E} \left[ \left( \frac{\tilde{X}_{i+1}^{(k)} - \tilde{X}_i^{(k)}}{\sqrt{k\Delta}} \right)^2 \middle| \mathcal{F}_t \right] = \sigma^2 + k\Delta a + k\Delta \tilde{\sigma} =: s_i^k$$

Notice that

$$\frac{S^2}{S^1} \xrightarrow{\Delta \rightarrow 0} \begin{cases} 1 & \text{if } \sigma \neq 0 \\ 2 & \text{if } \sigma = 0 \end{cases}$$

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Notice that

$$\frac{s^2}{s^1} \xrightarrow{\Delta \rightarrow 0} \begin{cases} 1 & \text{if } \sigma \neq 0 \\ 2 & \text{if } \sigma = 0 \end{cases} \Rightarrow 1 - \frac{\log \frac{s^2}{s^1}}{\log 2} \xrightarrow{\Delta \rightarrow 0} \begin{cases} 1 & \text{if } \sigma \neq 0 \\ 0 & \text{if } \sigma = 0 \end{cases}$$

# What happens if $\Delta$ is fixed?

Take the process:

$$dX_t = bdt + \sigma dW_t$$

**Fix:**  $\Delta = 0.01$ ,  $\sigma = 0.05$ ,  $b = 1$ ,  $\tilde{\sigma} = 0.01$

Add the random perturbation:

$$\tilde{X}_t^{(1)} = bdt + \sigma dW_t + \sqrt{\Delta}\tilde{\sigma}\tilde{W}_t,$$

$$\tilde{X}_t^{(2)} = bdt + \sigma dW_t + \sqrt{2\Delta}\tilde{\sigma}\tilde{W}_t$$

Notice that

$$\mathbb{E} \left[ \frac{S^2}{S^1} \middle| \mathcal{F}_t \right] \approx 1.8 \quad \Rightarrow \quad \mathbb{E} [\hat{R}_{T,\Delta} | \mathcal{F}_t] \approx 0.15$$

# What happens if $\Delta$ is fixed?

## Ongoing work

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### **Question 1: What can we actually infer in non-asymptotic setting?**

Our ultimate goal is to evaluate the following probability:

$$\mathbb{P} \left( \left| \widehat{R}_{T,\Delta} - r_0 \right| \geq \varepsilon \right) \quad (2)$$

### **Question 2: Choice of $\widetilde{\Sigma}$ .**

# Concentration inequality for 1d toy model

Consider a one-dimensional process with constant drift and diffusion coefficients:

$$dX_t = bdt + \sigma dW_t$$

Statistics:

$$S^k = 2\Delta \sum_{i=0}^N \left( \frac{\tilde{X}_{(2i+k)\Delta}^{(k)} - \tilde{X}_{2i\Delta}^{(k)}}{\sqrt{k\Delta}} \right)^2 \quad (3)$$

$S^k$  is difficult to treat, but luckily  $\sqrt{S^k}$  is a Lipschitz function of a normal vector!

$$\mathbb{P} \left( \sqrt{S^k} - \mathbb{E}[\sqrt{S^k}] \geq \varepsilon | \mathcal{F}_t \right) \leq e^{-\frac{\varepsilon^2}{2(\sigma^2 + 2\Delta\bar{\sigma}^2)}}$$



# Concentration inequality for 1d toy model

How do we pass from bounds on  $S^k$  to bounds on  $\hat{R}(T, \Delta)$ :

$$\begin{aligned} & \mathbb{P} \left( \frac{\mathbb{E}\sqrt{S^2} - \varepsilon}{\mathbb{E}\sqrt{S^1} + \varepsilon} \leq \frac{\sqrt{S^2}}{\sqrt{S^1}} \leq \frac{\mathbb{E}\sqrt{S^2} + \varepsilon}{\mathbb{E}\sqrt{S^1} - \varepsilon} \middle| \mathcal{F}_t \right) = \\ & \mathbb{P} \left( \log \frac{\mathbb{E}\sqrt{S^2} - \varepsilon}{\mathbb{E}\sqrt{S^1} + \varepsilon} \leq \log \frac{\sqrt{S^2}}{\sqrt{S^1}} \leq \log \frac{\mathbb{E}\sqrt{S^2} + \varepsilon}{\mathbb{E}\sqrt{S^1} - \varepsilon} \middle| \mathcal{F}_t \right) = \\ & \mathbb{P} \left( 1 - \frac{2}{\log 2} \log \frac{\mathbb{E}\sqrt{S^2} + \varepsilon}{\mathbb{E}\sqrt{S^1} - \varepsilon} \leq \hat{R}(T, \Delta) \leq 1 - \frac{2}{\log 2} \log \frac{\mathbb{E}\sqrt{S^2} - \varepsilon}{\mathbb{E}\sqrt{S^1} + \varepsilon} \middle| \mathcal{F}_t \right) \leq \\ & e^{-\frac{2\varepsilon^2}{(\sigma^2 + 2\Delta\bar{\sigma}^2)}} \end{aligned}$$

# Concentration inequality for 1d toy model

**How do we pass to confidence intervals:**  $\varepsilon := \sqrt{0.5(\sigma^2 + 2\Delta\tilde{\sigma}^2) \log \frac{1}{\beta}}$

Using an approximation of moments, we obtain:

$$\mathbb{P} \left( 1 - \frac{2}{\log 2} \log \frac{\sqrt{1 + \frac{2\Delta a^2}{\sigma^2 + 2\Delta\tilde{\sigma}^2}} + \sqrt{\frac{1}{2} \log \frac{1}{\beta}}}{\sqrt{1 + \frac{\Delta(a^2 - \tilde{\sigma}^2)}{\sigma^2 + 2\Delta\tilde{\sigma}^2}} - \sqrt{\frac{1}{2} \log \frac{1}{\beta}}} \leq \widehat{R}(T, \Delta) \leq \right. \\ \left. 1 - \frac{2}{\log 2} \log \frac{\sqrt{1 + \frac{2\Delta a^2}{\sigma^2 + 2\Delta\tilde{\sigma}^2}} - \sqrt{\frac{1}{2} \log \frac{1}{\beta}}}{\sqrt{1 + \frac{\Delta(a^2 - \tilde{\sigma}^2)}{\sigma^2 + 2\Delta\tilde{\sigma}^2}} + \sqrt{\frac{1}{2} \log \frac{1}{\beta}}} \middle| \mathcal{F}_t \right) \geq 1 - \beta$$

# Non-asymptotic setting: concentration inequalities

Recall main statistics:

$$S_T^k = 2d\Delta \sum_{i=0}^{N-1} \det \left( \begin{array}{ccc} \frac{\tilde{X}_{2id+k}^{1,(k)} - \tilde{X}_{2id}^{1,(k)}}{\sqrt{k\Delta}} & \frac{\tilde{X}_{2id+2k}^{1,(k)} - \tilde{X}_{2id+k}^{1,(k)}}{\sqrt{k\Delta}} & \cdots & \frac{\tilde{X}_{2id+2d}^{1,(k)} - \tilde{X}_{2id+kd-k}^{1,(k)}}{\sqrt{k\Delta}} \\ \frac{\tilde{X}_{2id+k}^{2,(k)} - \tilde{X}_{2id}^{2,(k)}}{\sqrt{k\Delta}} & \frac{\tilde{X}_{2id+2k}^{2,(k)} - \tilde{X}_{2id+k}^{2,(k)}}{\sqrt{k\Delta}} & \cdots & \frac{\tilde{X}_{2id+2d}^{2,(k)} - \tilde{X}_{2id+kd-k}^{2,(k)}}{\sqrt{k\Delta}} \\ \cdots & \cdots & \cdots & \cdots \\ \frac{\tilde{X}_{2id+k}^{d,(k)} - \tilde{X}_{2id}^{d,(k)}}{\sqrt{k\Delta}} & \frac{\tilde{X}_{2id+2k}^{d,(k)} - \tilde{X}_{2id+k}^{d,(k)}}{\sqrt{k\Delta}} & \cdots & \frac{\tilde{X}_{2id+2d}^{d,(k)} - \tilde{X}_{2id+kd-k}^{d,(k)}}{\sqrt{k\Delta}} \end{array} \right)^2$$

# General case: lower bound

## Lemma (Lower bound for $S^k$ )

$$\mathbb{P}(S^k - \mathbb{E}[S^k] \leq -\varepsilon) \leq \exp\left(\frac{-\varepsilon^2}{2Nv}\right),$$

where  $v = \sup_i \mathbb{E}[(s_i^k)^2]$ .

**Good news:** lower bound is sub-gaussian, because  $S^k$  is positive!

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**Good news:** lower bound is sub-gaussian, because  $S^k$  is positive!

**Bad news:** ...it depends on the expression of moments, which are difficult to compute!

# General case: upper bound

**Key:** Hadamard inequality

$$\left( \sum_{i=0}^N s_i^k \right)^{\frac{1}{2d}} \leq \left( \sum_{i=0}^N \left( \sup_t \left\| \frac{X_{t+k\Delta}^{(k)} - X_t^{(k)}}{\sqrt{k\Delta}} \right\| \right)^{2d} \right)^{\frac{1}{2d}} \leq \sum_{i=0}^N \sup_t \left\| \frac{X_{t+k\Delta}^{(k)} - X_t^{(k)}}{\sqrt{k\Delta}} \right\|$$

Lemma (Upper bound for  $S^k$ )

$$\mathbb{P} \left( (S^k)^{\frac{1}{2d}} - \mathbb{E}[(S^k)^{\frac{1}{2d}}] \geq \varepsilon \right) \leq \exp \left( \frac{-\varepsilon^2}{2Nv} \right),$$

where  $v = \sup_t (\Sigma + 2\Delta\tilde{\Sigma})$ .

# Pasting it all together....

$$\begin{aligned} \mathbb{P} \left( \frac{1}{2d} \ln \frac{\mathbb{E}[S^2] - \varepsilon}{\mathbb{E}[S^1]^{\frac{1}{2d}} + \varepsilon} \geq \ln \frac{S^2}{S^1} \geq 2d \ln \frac{\mathbb{E}[S^2]^{\frac{1}{2d}} + \varepsilon}{\mathbb{E}[S^1] - \varepsilon} \middle| \mathcal{F}_t \right) = \\ \mathbb{P} \left( d - \frac{2d}{\ln 2} \ln \frac{\mathbb{E}[S^2]^{\frac{1}{2d}} + \varepsilon}{\mathbb{E}[S^1] - \varepsilon} \geq \hat{R}(T, \Delta) \geq d - \frac{1}{2d \ln 2} \ln \frac{\mathbb{E}[S^2] - \varepsilon}{\mathbb{E}[S^1]^{\frac{1}{2d}} + \varepsilon} \middle| \mathcal{F}_t \right) \\ \leq \exp \left( -\frac{2\varepsilon^2}{N\nu} \right) \end{aligned}$$

# Open questions: part I

- ▶ Is it the sharpest bound? Can we get rid of  $N$ ?
- ▶ **How to evaluate**  $\mathbb{E}[S^k]$ ?



# Open questions: part I

- ▶ Is it the sharpest bound? Can we get rid of  $N$ ?
- ▶ **How to evaluate  $\mathbb{E}[S^k]$ ?**
  - **If the rows of the matrix are independent:** statistics  $s_i^k$  behave as a product of non-centered chi-squared variables
  - **If they are not:** no general result known [*to the authors*].

## ...some flavour of it

The upper and the lower bound for the first moment of  $S^k$  in the case of **independent rows** is the following:

$$\begin{aligned} & \sum_{i=0}^N \left[ d! \prod_{i=1}^d \min_{j=1, \dots, d} \sigma_{\phi_{i,j}^k}^2 + \right. \\ & \left. k\Delta(d-1)! \left( \sum_{l,j=0}^d \mu_{\phi_{i,j}^k} \right)^2 \prod_{l=1, l \neq j}^{d-1} \min_{j=1, \dots, d} \sigma_{\phi_{i,j}^k}^2 + O(\Delta^2) \right] \leq \mathbb{E} [S^k | \mathcal{F}] \\ & \leq \sum_{i=0}^N \left[ d! \prod_{i=1}^d \max_{j=1, \dots, d} \sigma_{\phi_{i,j}^k}^2 + \right. \\ & \left. k\Delta(d-1)! \left( \sum_{l,j=0}^d \mu_{\phi_{i,j}^k} \right)^2 \prod_{l=1, l \neq j}^{d-1} \max_{j=1, \dots, d} \sigma_{\phi_{i,j}^k}^2 + O(\Delta^2) \right], \end{aligned}$$

# Open questions: part II

- ▶ ... can we potentially tell the difference between the "true" (of order  $\sqrt{\Delta}$ ) and the propagated (of order  $\Delta^{\frac{3}{2}}$ ) noise in a given variable?
- ▶ ... what is the best empirical choice of  $\tilde{\Sigma}$ ? **Does the optimal  $\tilde{\Sigma}$  even exist?**

# Numerical experiments: general setting

1. Generate 1000 trajectories with  $\Delta = 1e - 5$ ,  $T = 10$ , using 1.5 strong order scheme (see Kloeden et al. (2003))
2. Subsample the data with a bigger  $\Delta$
3. Compute test statistics  $\hat{R}_{T,\Delta}$  and  $V_{T,\Delta}$
4. Test the "true" and a "wrong" hypothesis for a nominal level  $\alpha = 0.05$
5. Report the results

# Example 1: FitzHugh-Nagumo model

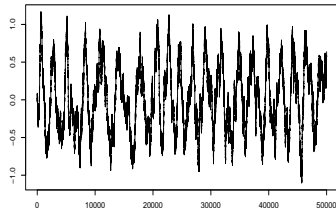
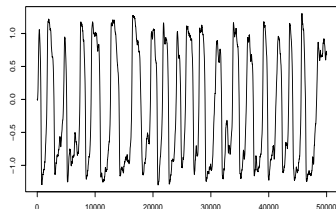
The behaviour of the neuron is defined through:

$$\begin{cases} dX_t = \frac{1}{\varepsilon}(X_t - X_t^3 - Y_t - s)dt + \sigma_1 dW_t^1 \\ dY_t = (\gamma X_t - Y_t + \beta)dt + \sigma_2 dW_t^2 \end{cases}$$

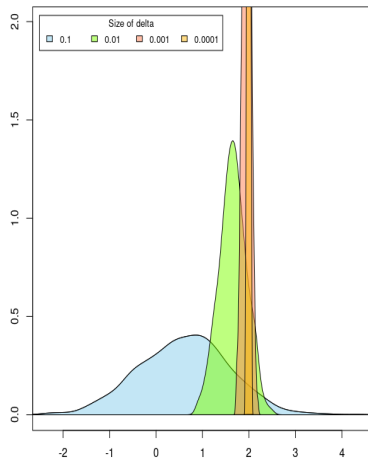
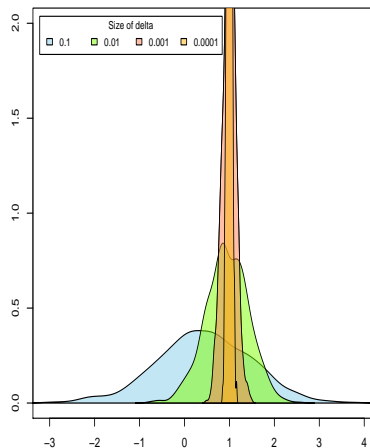
- ▶  $X_t$  – membrane potential
- ▶  $Y_t$  – recovery variable
- ▶  $s$  – magnitude of the stimulus current

**Parameters** used in simulations:

$$\varepsilon = 0.1, \beta = 0.3, \gamma = 1.5, s = 0.01.$$

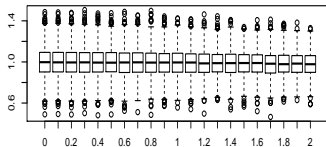


# Numerical performance: FitzHugh-Nagumo model

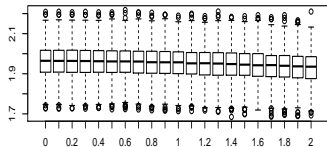


# Numerical performance: FitzHugh-Nagumo model

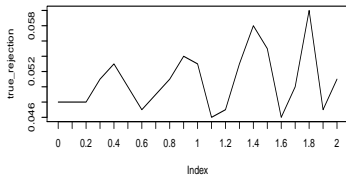
Estimator boxplot



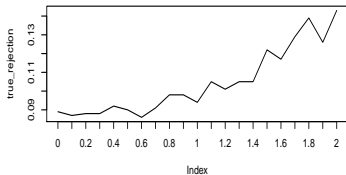
Estimator boxplot



rejection of true hypothesis



rejection of true hypothesis



# Literature I

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Thank you for your attention!